

# Spatial Statistics: Beyond a Textbook

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# Outline

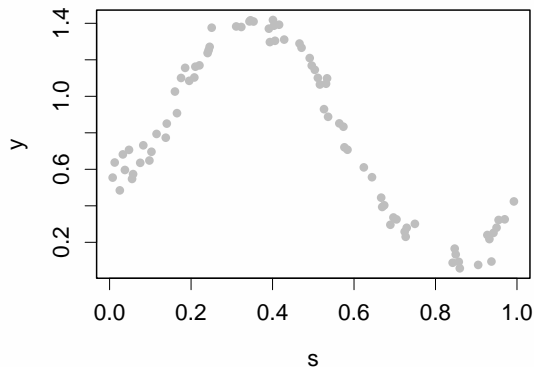
- Gaussian process models
- A 1D example
- The precision matrix
- Approximation for large data - basis functions
- Change of support
- Non-Gaussian observations

Waldo Tobler:

*“everything is related to everything else, but near things are more related than distant things.”*

... but the distant things become less important given the near things ...

## Some data



Goal: Inference for the maximum – where is it? How high? What is the uncertainty?

# Gaussian Process model

$g(s)$  Gaussian process

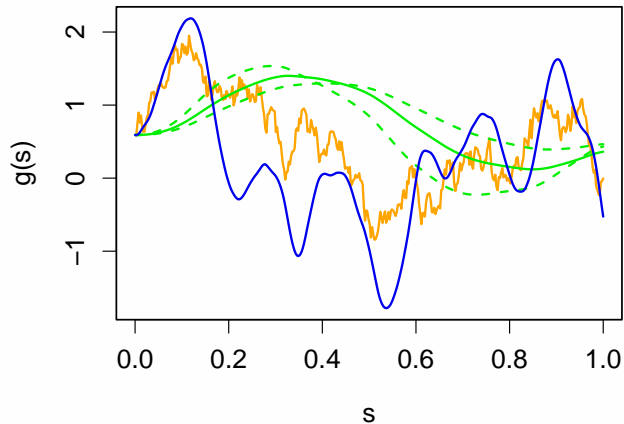
- $E(g(s)) = 0$
- $Cov(g(s), g(s')) = k(s, s')$

Some examples of stationary  $k(., .)$

$d = \|s - s'\|/\alpha$  Distance scaled by a parameter.

- Exponential:  $k(s, s') = e^{-d}$
- Gaussian:  $k(s, s') = e^{-d^2}$
- Matern ( $\nu = 2.5$ ):  $k(s, s') = (1 + d + (1/3)d^2)e^{-d}$
- Wendland:  $k(s, s') = \begin{cases} (1 - d)^6(35d^2 + 18d + 3)/3 & d \leq 0 \\ 0 & d > 0 \end{cases}$

# Gaussian Process realizations



Exponential, Matern, Wendland

# The data

$$Y_i = g(s_i) + e_i$$

$\{e_i\}$  Independent  $N(0, \tau^2)$

Assume that  $\text{VAR}(g(s_i)) = \sigma^2$

The Big Three:

- Correlation range,  $\alpha$
- Variance of process,  $\sigma^2$
- Variance of measurement error,  $\tau^2$

# What we do

Goal: estimate/infer  $g$  from  $\mathbf{z}$  ( but we will need the Big Three to do it).

The program:

- Estimate the big three by maximum likelihood (or Bayes).
- Sample from the conditional distribution of  $g$  given the observations.

# Kriging – finding $g$

Recall

$$\mathbf{z} = \mathbf{g} + \mathbf{e}$$

$$\mathbf{e} = \begin{pmatrix} g(s_1) \\ g(s_2) \\ \vdots \\ g(s_n) \end{pmatrix}$$

Assumptions

- $[\mathbf{z}|\mathbf{g}]$  is  $MN(\mathbf{g}, \tau^2 I)$
- $[\mathbf{g}]$  is  $MN(0, \sigma^2 K)$
- $[\mathbf{z}|\mathbf{g}][\mathbf{g}] = [\mathbf{z}, \mathbf{g}]$  is jointly multivariate normal.

Note: we can fill in  $K$  for any set of locations because we have the covariance function.



## More on Kriging

Estimate is the conditional expectation of  $g(s)$  given  $\mathbf{z}$ .

$$\hat{g}(s) = E[g(s)|\mathbf{z}] = \sum_{i=1}^n k(s, s_i)\mathbf{c}_i$$

with  $\mathbf{c}$  is the solution to the linear system

$$\text{COV}(\mathbf{z}, \mathbf{z})\mathbf{c} = \mathbf{z}$$

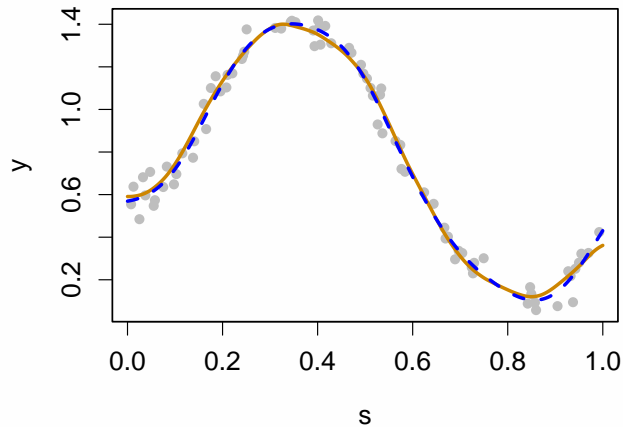
$$(\sigma^2 K(\alpha) + \tau^2 I)\mathbf{c} = \mathbf{z}$$

Prediction at the data is a smoother:

$$\hat{h} = K(\alpha)(K(\alpha) + \lambda I)^{-1}\mathbf{z}$$

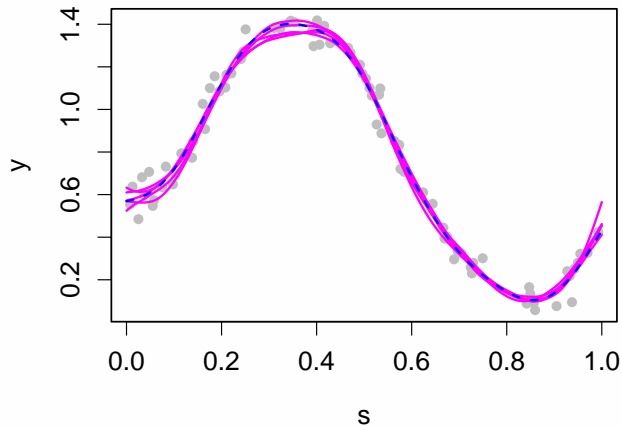
$$\lambda = \tau^2/\sigma^2$$

# Estimate

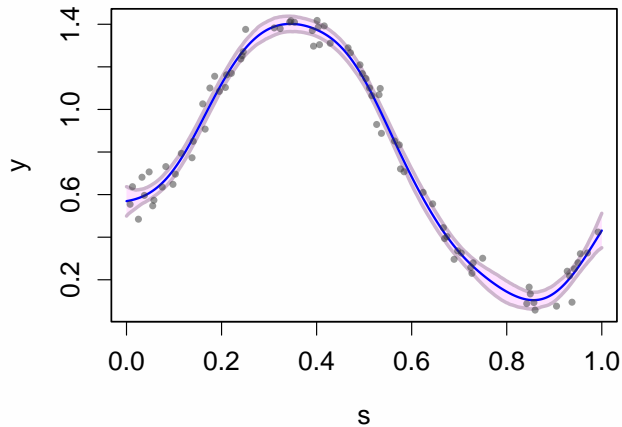


true curve

## 5 Draws from $[g|z]$



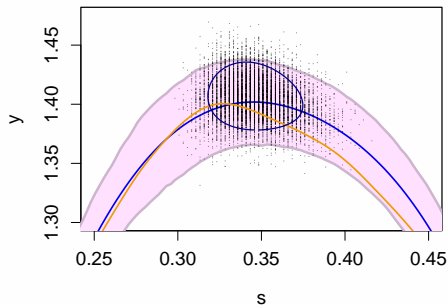
## 95% envelope



## Inference for the maximum

$\approx 95\%$  confidence region for maximum  $X$  and  $Y$

Points are 1000 realizations of the maxima from conditional samples



True curve, Estimate

We are 95% confident the maximum of  $g$  and its location are in this region.

## What about the big three?

Integrate out the latent Gaussian process.

$$\int [\mathbf{z}, \mathbf{g} | \sigma^2, \tau^2, \alpha] d\mathbf{g} = [\mathbf{z} | \sigma^2, \tau^2, \alpha] \sim MN(0, \sigma^2 K(\alpha) + \tau^2 I)$$

log likelihood is

$$-\frac{1}{2} \mathbf{z}^T (\sigma^2 K(\alpha) + \tau^2 I)^{-1} \mathbf{z} - \frac{1}{2} \log(|\sigma^2 K(\alpha) + \tau^2 I|) + \text{constant}$$

Can concentrate over  $\lambda = \tau^2 / \sigma^2$  and  $\alpha$  for a 2D optimization.

For a few thousand locations (in `fields` package)

```
obj <- spatialProcess(s,z, smoothness=2.5)
```

Exact linear algebra increases as  $O(n^3)$  with  $n$  the number of locations.

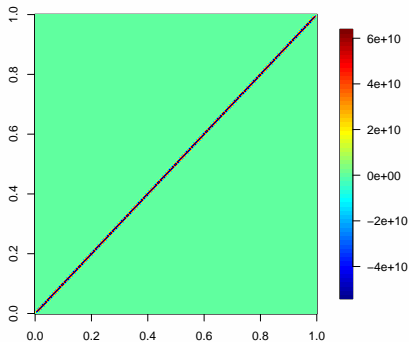
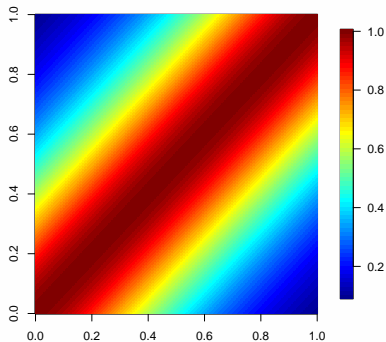
$n = 3500$ , takes 30 seconds  $\rightarrow n = 8 * 3500$ , takes  $\approx 4$  hours

# More about covariance matrices

Covariance used in the example

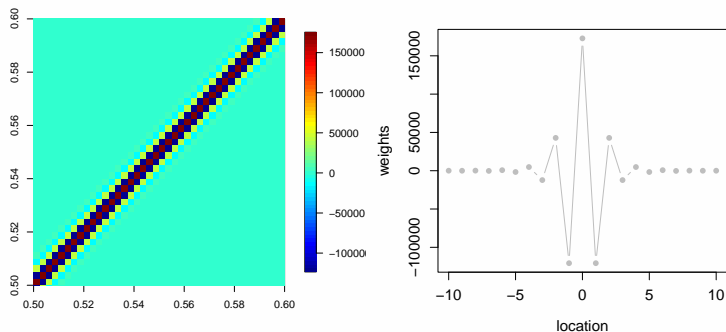
Covariance matrix (Matern  $\nu = 2.5$ )

and its inverse



# Spatial AR

Symmetric Inverse square root of covariance matrix



If  $\text{cov}(\mathbf{g}) = \Sigma$  then  $\Sigma^{-1/2}\mathbf{g}$  has an identity covariance matrix.

i.e.

$$\text{cov}(\Sigma^{-1/2}\mathbf{g}) = \Sigma^{-1/2}(\Sigma)\Sigma^{-1/2} = I$$

- $\Sigma^{-1/2}$  a recipe for decorrelation.



# Fixed Rank Kriging (FRK): Basic Idea

- Model:

$$\mathbf{z}_i = \mathbf{X}_i\beta + \mathbf{g}(s_i) + \epsilon_i$$

- The spatial process  $g(\cdot)$  is approximated by a linear combination of  $L$  fixed basis functions and random coefficients.

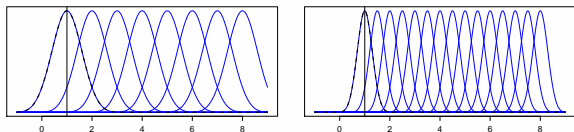
$$g(s) = \psi_1(s)c_1 + \psi_2(s)c_2 + \dots + \psi_L(s)c_L$$

- Assume that  $\mathbf{c}$  is multivariate normal, mean zero and covariance matrix,  $Q^{-1}$ .
- Sherman-Morrison-Woodbury formula reduces the computations to linear algebra in  $Q$  ( $L \times L$ ).  
If  $Q$  is dense then  $L$  small is helpful.  
If  $Q$  is sparse then  $L$  can be larger than the number of observations.
- $Q = B^T B$  then the rows of  $B$  are a recipe for transforming the coefficients to independent  $N(0, 1)$

## More on basis functions and coefficients

$$g(s) = \psi_1(s)c_1 + \psi_2(s)c_2 + \dots + \psi_L(s)c_L$$

In practice,  $\{\psi_j\}$  have different resolutions and are zero beyond a fixed range.



Model for the coefficients,  $j^{\text{th}}$  row of  $B$ :

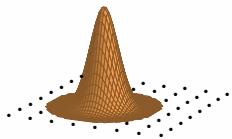
$$\dots \quad 0 \quad 0 \quad -1 \quad a_j \quad -1 \quad 0 \quad 0 \quad \dots$$

The recipe for decorrelation:

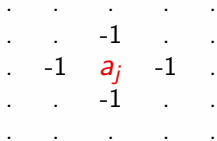
$$a_j \mathbf{c}_j - \mathbf{c}_{j-1} - \mathbf{c}_{j+1} \sim N(0, 1)$$

Note  $a_j > 2$  for stationarity

# LatticeKrig 2 D Implementation



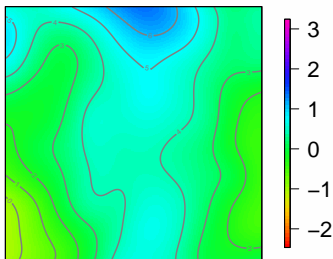
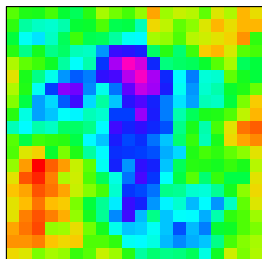
A 2-d basis function



Need  $a_j > 4$

$16^2$  Coefficients

Expanding with basis functions

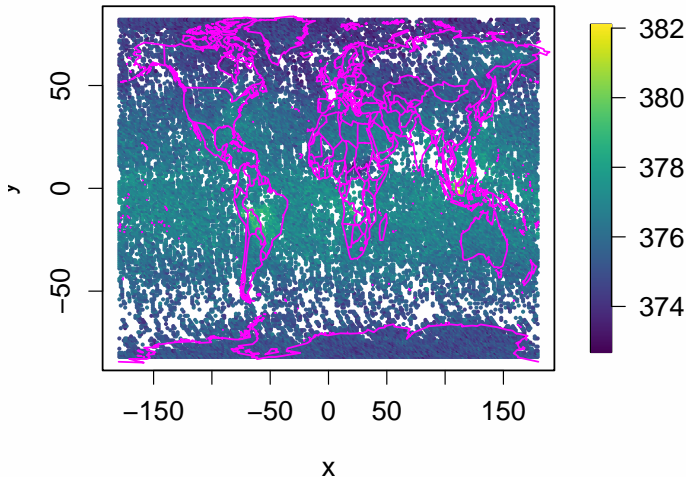


$$\mathbf{c}_j \rightarrow \sum \psi_j(x) \mathbf{c}_j$$

# LatticeKrig Demo in R

CO2 data set  $\approx$  27K observations in PPM.

**Total column CO2**



## Change of support

Observations of area averages

$$g(s) = \psi_1(s)c_1 + \dots + \psi_L(s)c_L$$

$$\int_A g(s)ds = \left( \int_A \psi_1(s)ds \right) c_1 + \dots + \left( \int_A \psi_L(s)ds \right) c_L$$

- Observations:

$z_i$  average over region  $A_i$

$$\mathbf{z} = \mathbf{W}\mathbf{c} + \mathbf{e}$$

$W_{i,j}$  integral of region  $i$  over basis function,  $j$ .

- Spatial process:

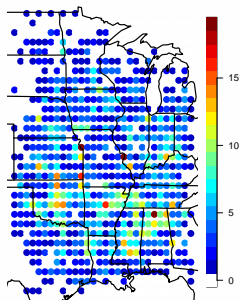
$g(s)$  follows *same* fixed rank Kriging model,  $\mathbf{c} \sim MN(0, Q^{-1})$

# Non Gaussian observations

Poisson counts with a spatial intensity

- $\mathbf{z}_i \sim \text{Poisson}(\gamma_i)$
- $\log(\gamma(\mathbf{s})) \sim$  Gaussian process with  $\sigma^2$  and  $\alpha$  and mean  $X\beta$ .
- Prior on  $\sigma^2$ ,  $\alpha$  and  $\beta$

*US Tornado counts 1950-2020 (F level 3 or higher)*



$X =$  lon, lat, and elevation

## As a Hierarchical Model

Observational model:  $[\mathbf{z}_i | \mathbf{g}, \beta, \alpha, \sigma^2] \sim \text{Poisson}(e^{X_i \beta + g(s_i)})$

Process:  $[\mathbf{g} | \beta, \alpha, \sigma^2] \sim N(X\beta, \sigma^2 K(\alpha))$

Priors:  $[\beta, \alpha, \sigma^2]$

Key technical step is to "integrate out"  $\mathbf{g}$

$$[\mathbf{z}_i | \beta, \alpha, \sigma^2] = \int [\mathbf{z}_i | \mathbf{g}, \beta, \alpha, \sigma^2] d\mathbf{g}$$

- Needed to find MLEs or Bayesian inference.
- Use a Laplace approximation for this step – this is the basis of the INLA framework.

## Some details for $g(s)$

Poisson case:

$$[\mathbf{g}, \alpha, \sigma^2, \beta | \mathbf{z}] \propto [\mathbf{z} | \beta, \mathbf{g}] [\mathbf{g} | \alpha, \sigma^2]$$

$$\propto \left[ \prod_{i=1}^n (\gamma_i^{z_i}) e^{-\gamma_i} \right] \left[ \frac{e^{-(\mathbf{g} - X\beta)^T (\sigma^2 K(\alpha))^{-1} (\mathbf{g} - X\beta) / 2}}{|\sigma^2 K(\alpha)|^{1/2}} \right]$$

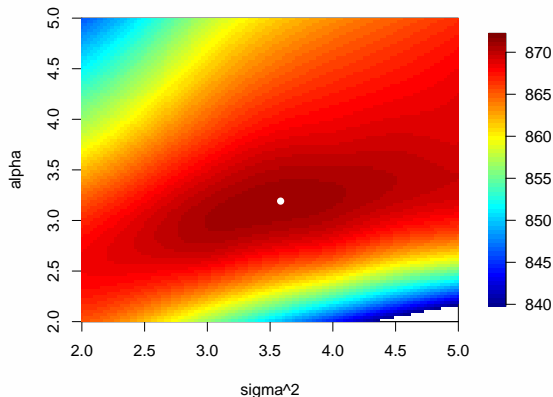
and  $\gamma_i = e^{x_i^T \beta + g_i}$  or  $\log(\gamma) = X\beta + \mathbf{g}$

- For fixed parameters can maximize over  $\mathbf{g}$  – this is a crucial computation giving –  $\hat{\mathbf{g}}$
- Note that  $\hat{\mathbf{g}}$  depends on the data and other parameters.
- Expand the log of the likelihood in a Taylor's series about  $\hat{\mathbf{g}}$  giving a multivariate normal approximation.

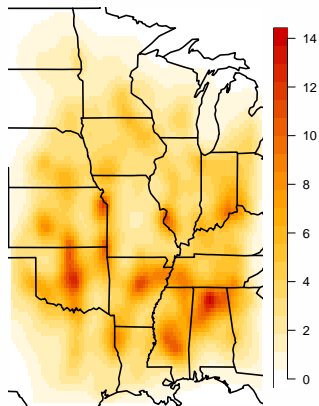


# Intensity of tornados

Laplace approximation to log likelihood surface for  $\sigma^2$  and  $\alpha$



## Intensity of tornados



Predicted intensity surface – is this biased towards population centers?

## Conclusions

- Spatial analysis tied to multivariate normal that describes dependence
- Conditional sampling from the spatial model is a useful for inference.
- Dependence can be introduced into a basis function/ coefficient way to handle large data and complicated observation types.

Methods in the `fields` and `LatticeKrig` R packages.

# Thank you!

